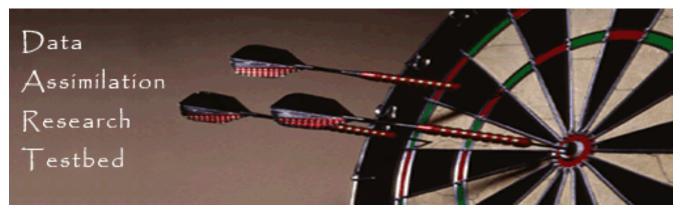
Ensemble Data Assimilation and Uncertainty Quantification



Jeff Anderson National Center for Atmospheric Research

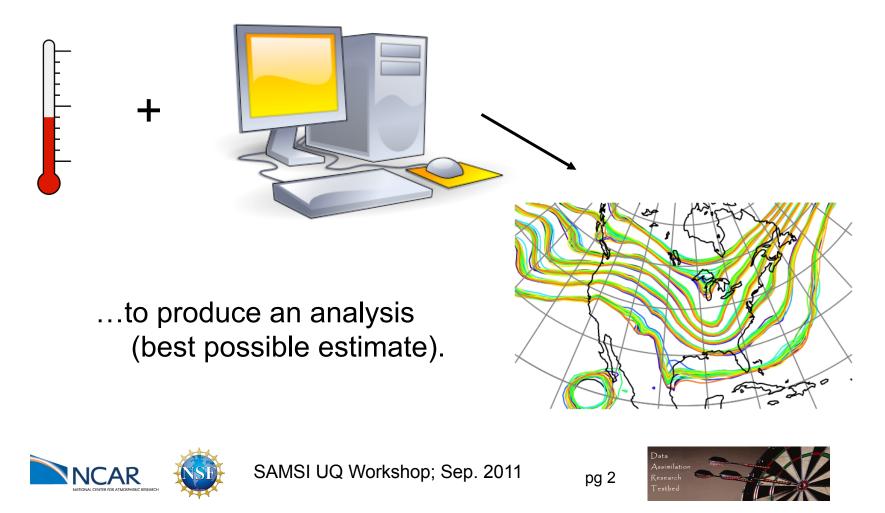


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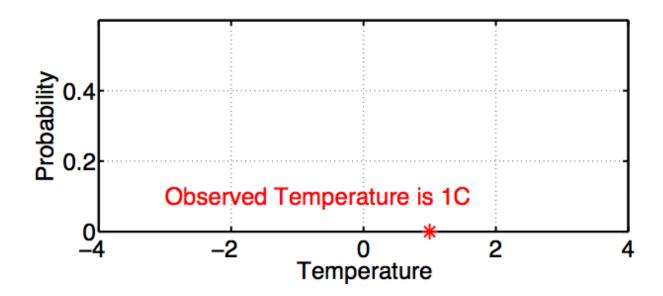


What is Data Assimilation?

Observations combined with a Model forecast...



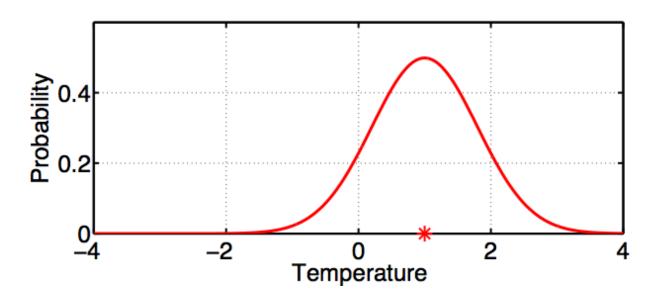
An observation has a value (*),







An observation has a value (*),

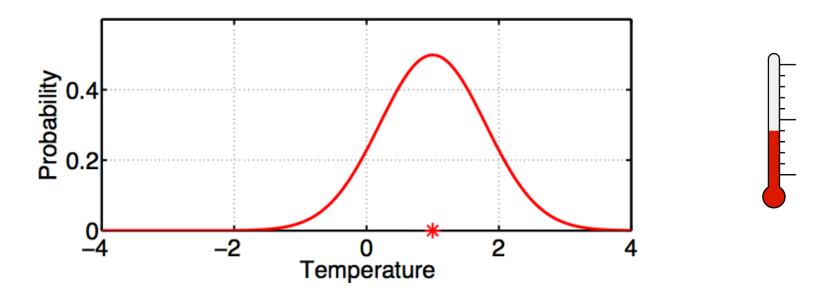


and an error distribution (red curve) that is associated with the instrument.





Thermometer outside measures 1C.



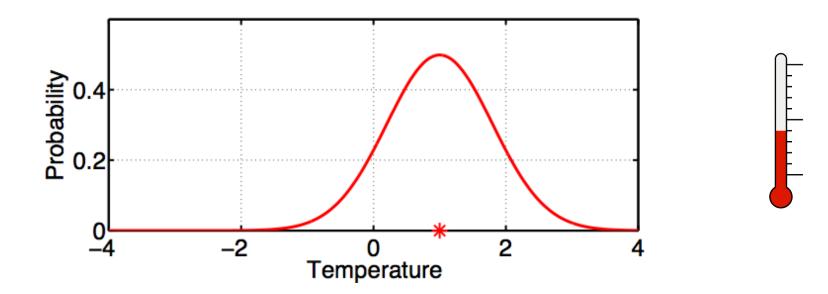
Instrument builder says thermometer is unbiased with +/- 0.8C gaussian error.



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Thermometer outside measures 1C.

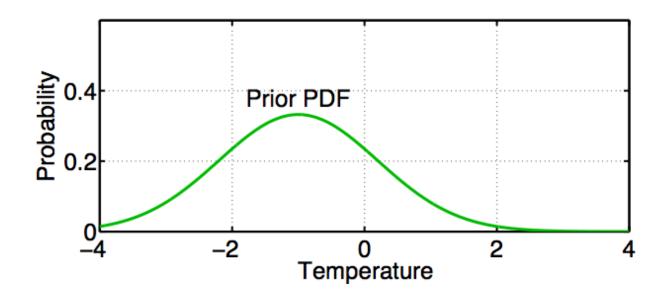


The red plot is $P(T \mid T_o)$, probability of temperature given that T_o was observed.





We also have a prior estimate of temperature.



The green curve is P(T | C); probability of temperature given all available prior information *C*.



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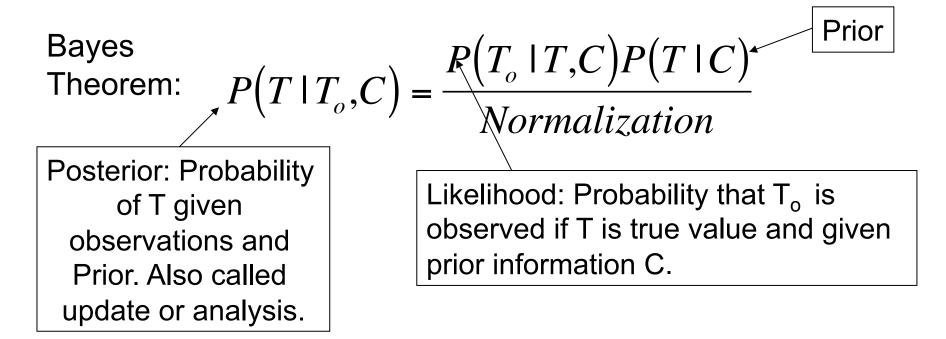


Prior information C can include:

- 1. Observations of things besides T;
- 2. Model forecast made using observations at earlier times;
- 3. A priori physical constraints (T > -273.15C);
- 4. Climatological constraints (-30C < T < 40C).

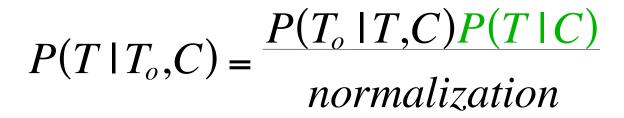


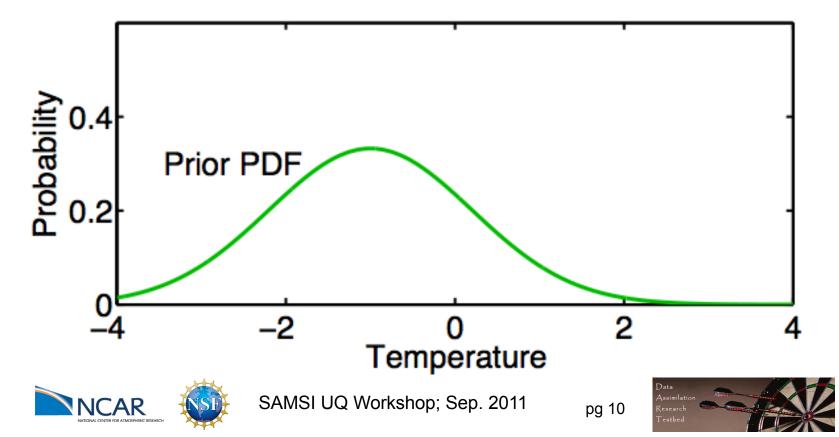


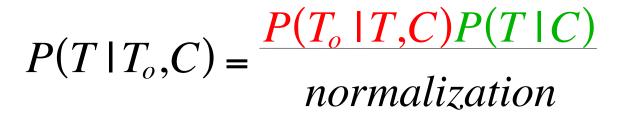


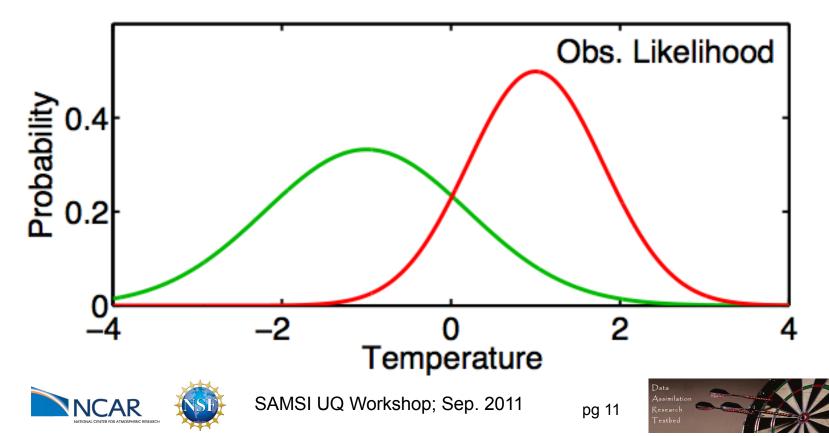




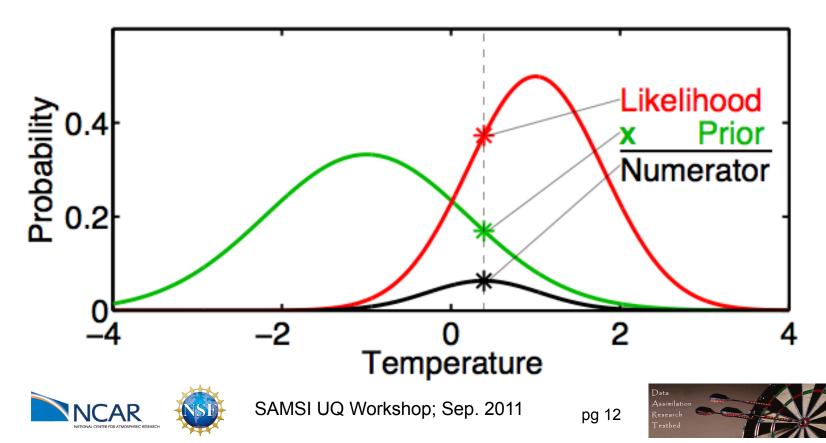




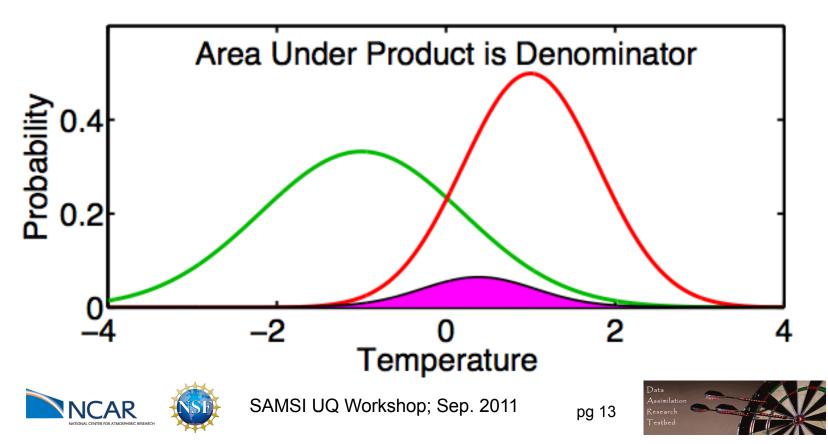


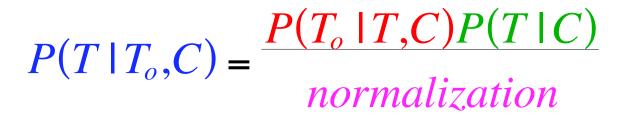


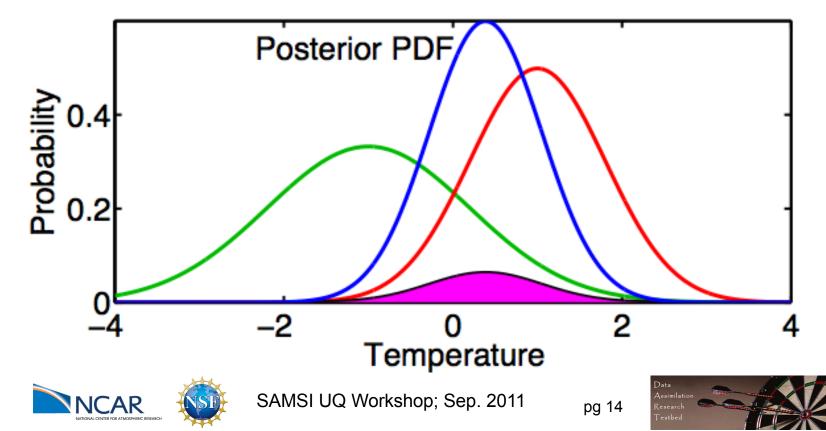




$$P(T \mid T_o, C) = \frac{P(T_o \mid T, C)P(T \mid C)}{normalization}$$







Consistent Color Scheme Throughout Tutorial

Green = Prior

Red = Observation

Blue = Posterior

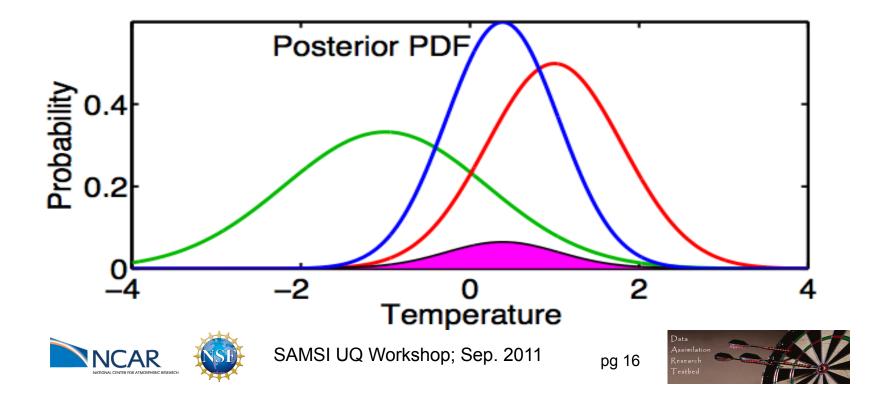
Black = Truth





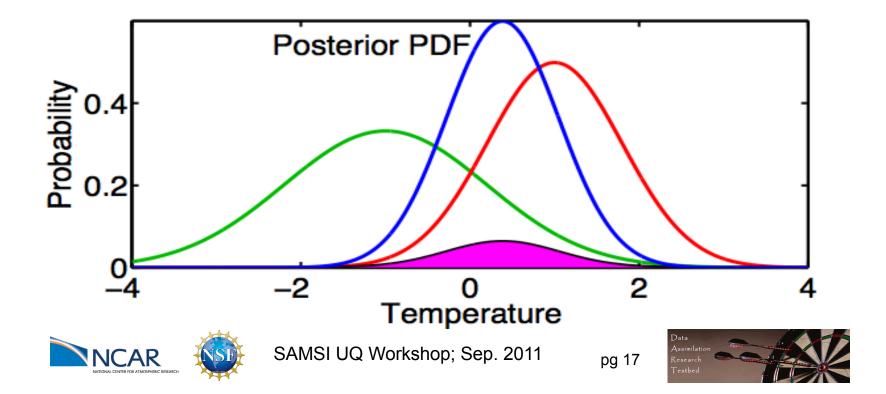
$$P(T \mid T_o, C) = \frac{P(T_o \mid T, C)P(T \mid C)}{normalization}$$

Generally no analytic solution for Posterior.



$$P(T | T_o, C) = \frac{P(T_o | T, C)P(T | C)}{normalization}$$

Gaussian Prior and Likelihood -> Gaussian Posterior



Combining the Prior Estimate and Observation For Gaussian prior and likelihood...

- Prior $P(T | C) = Normal(T_p, \sigma_p)$ Likelihood $P(T_o | T, C) = Normal(T_o, \sigma_o)$
- Then, Posterior $P(T | T_o, C) = Normal(T_u, \sigma_u)$

$$\sigma_u = \sqrt{\left(\sigma_p^{-2} + \sigma_o^{-2}\right)^{-1}}$$

With

$$T_u = \sigma_u^2 \left[\sigma_p^{-2} T_p + \sigma_o^{-2} T_o \right]$$





- 1. Suppose we have a linear forecast model L
 - A. If temperature at time $t_1 = T_{1,}$ then temperature at $t_2 = t_1 + \Delta t$ is $T_2 = L(T_1)$
 - B. Example: $T_2 = T_1 + \Delta t T_1$





- 1. Suppose we have a linear forecast model L.
 - A. If temperature at time $t_1 = T_{1,}$ then temperature at $t_2 = t_1 + \Delta t$ is $T_2 = L(T_1)$.
 - B. Example: $T_2 = T_1 + \Delta t T_1$.
- 2. If posterior estimate at time t_1 is *Normal*($T_{u,1}, \sigma_{u,1}$) then prior at t_2 is *Normal*($T_{p,2}, \sigma_{p,2}$).

$$\mathsf{T}_{\mathsf{p},2} = \mathsf{T}_{\mathsf{u},1}, + \Delta \mathsf{t} \mathsf{T}_{\mathsf{u},1}$$

$$\sigma_{p,2} = (\Delta t + 1) \sigma_{u,1}$$





- 1. Suppose we have a linear forecast model L.
 - A. If temperature at time $t_1 = T_{1,}$ then temperature at $t_2 = t_1 + \Delta t$ is $T_2 = L(T_1)$.
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- 3. Given an observation at t_2 with distribution *Normal*(t_0, σ_0) the likelihood is also *Normal*(t_0, σ_0).





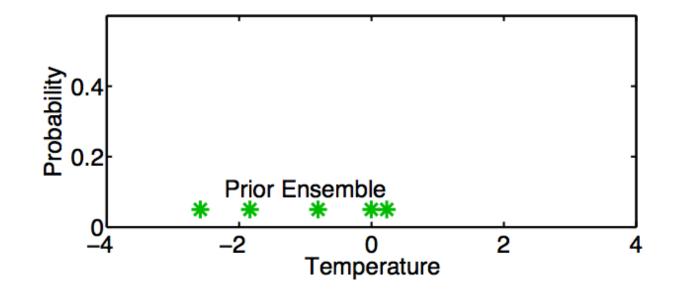
- 1. Suppose we have a linear forecast model L.
 - A. If temperature at time $t_1 = T_{1,}$ then temperature at $t_2 = t_1 + \Delta t$ is $T_2 = L(T_1)$.
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- 3. Given an observation at t_2 with distribution *Normal*(t_0, σ_0) the likelihood is also *Normal*(t_0, σ_0).
- 4. The posterior at t₂ is *Normal*($T_{u,2}$, $\sigma_{u,2}$) where $T_{u,2}$ and $\sigma_{u,2}$ come from page 18.





A One-Dimensional Ensemble Kalman Filter

Represent a prior pdf by a sample (ensemble) of N values:

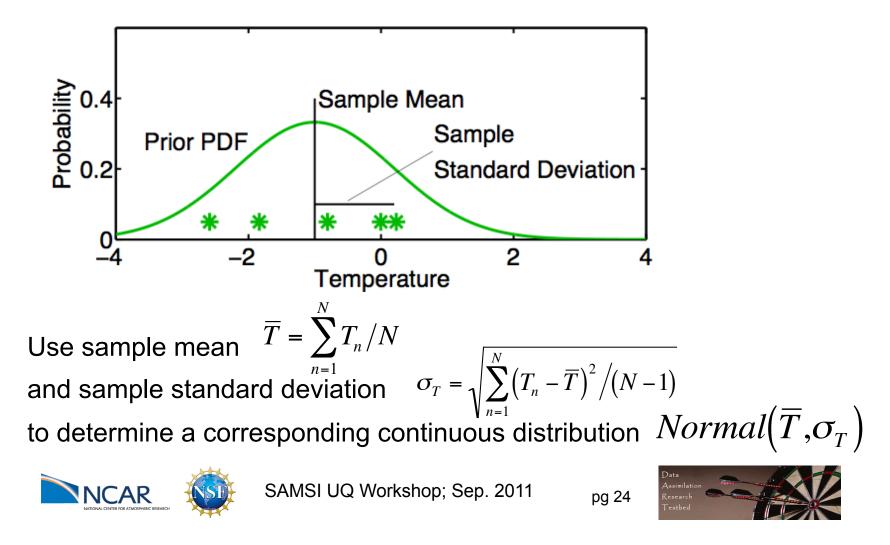




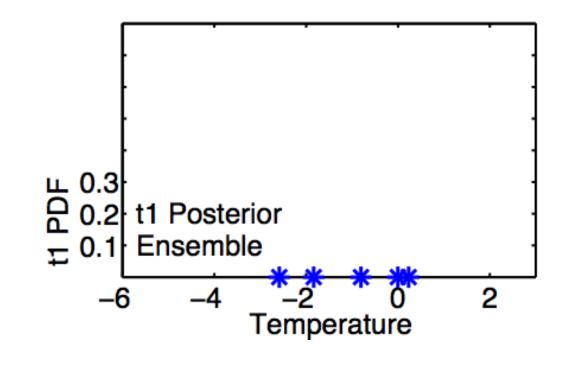


A One-Dimensional Ensemble Kalman Filter

Represent a prior pdf by a sample (ensemble) of N values:



If posterior ensemble at time t_1 is $T_{1,n}$, n = 1, ..., N

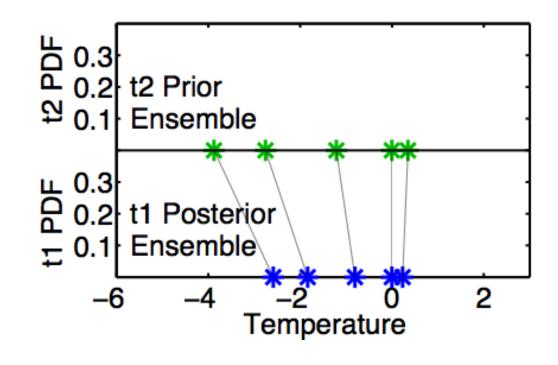




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If posterior ensemble at time t_1 is $T_{1,n}$, n = 1, ..., N, advance each member to time t_2 with model, $T_{2,n} = L(T_{1,n})$, n = 1, ..., N.

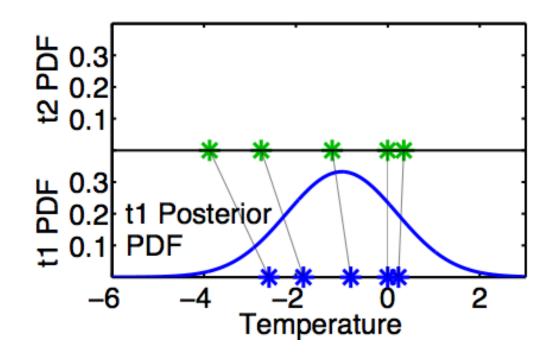




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Same as advancing continuous pdf at time $t_1 \dots$

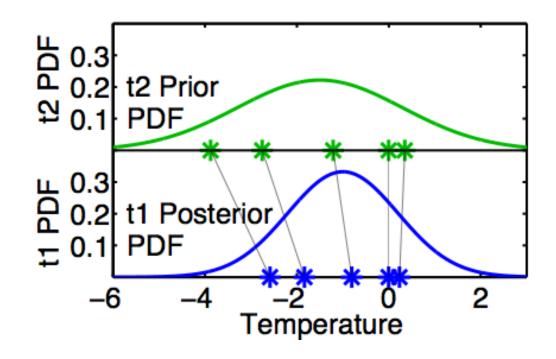




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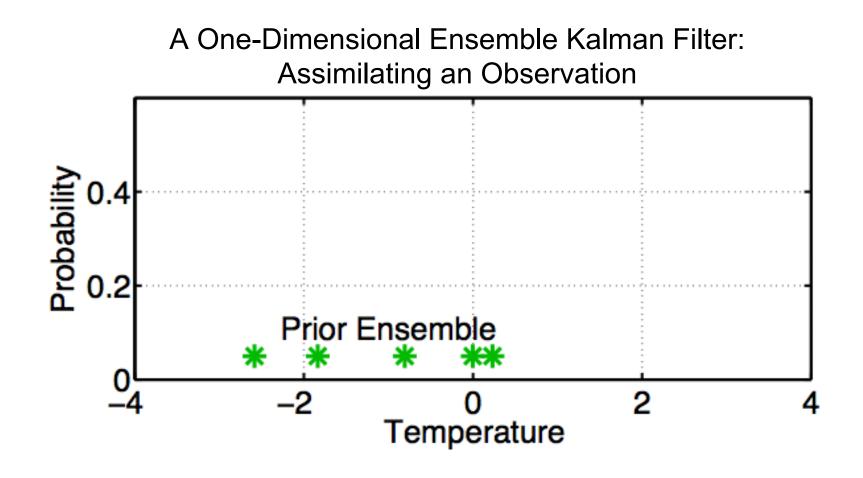
Same as advancing continuous pdf at time t_1 to time t_2 with model L.





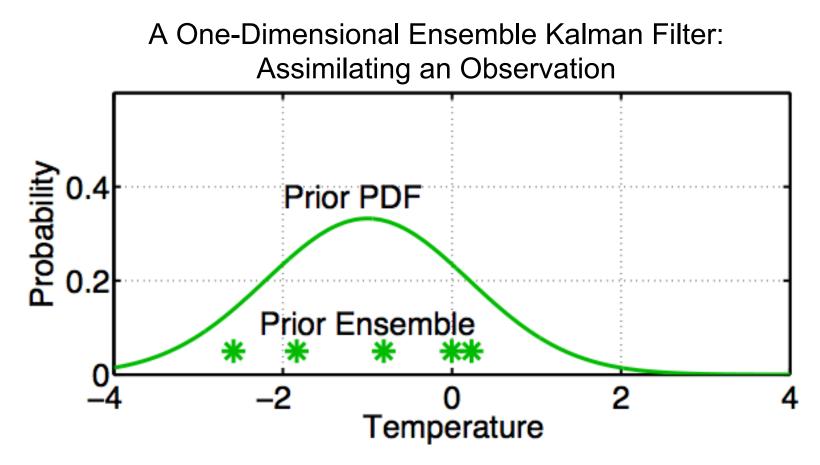
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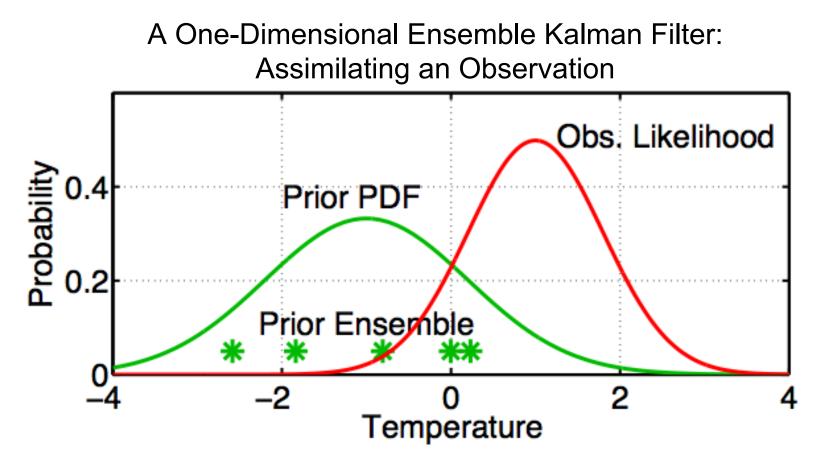




Fit a Gaussian to the sample.



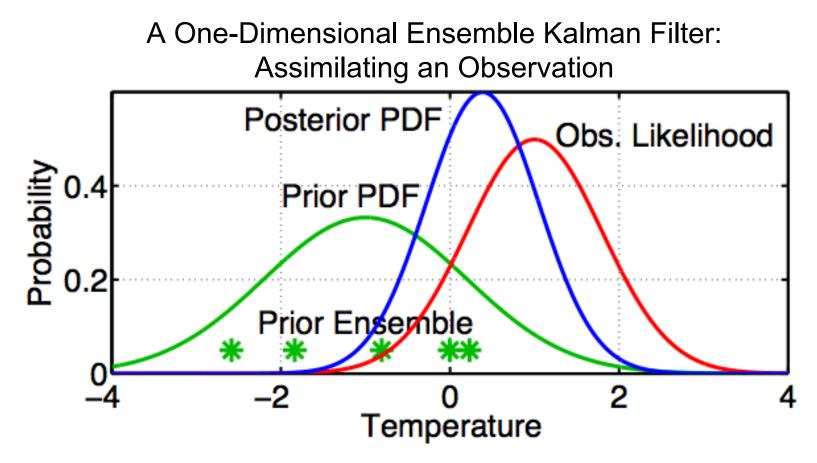




Get the observation likelihood.



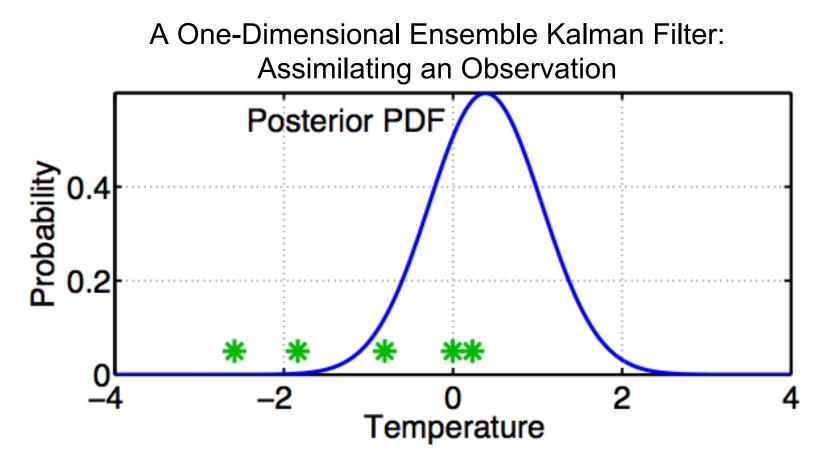




Compute the continuous posterior PDF.



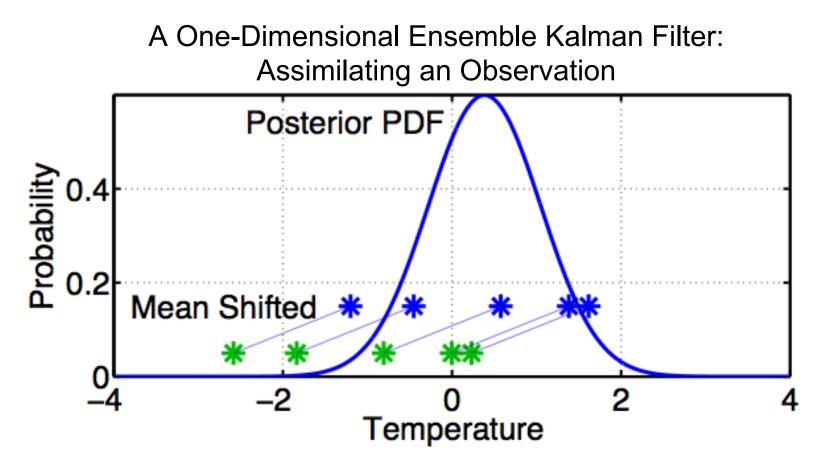




Use a deterministic algorithm to 'adjust' the ensemble.



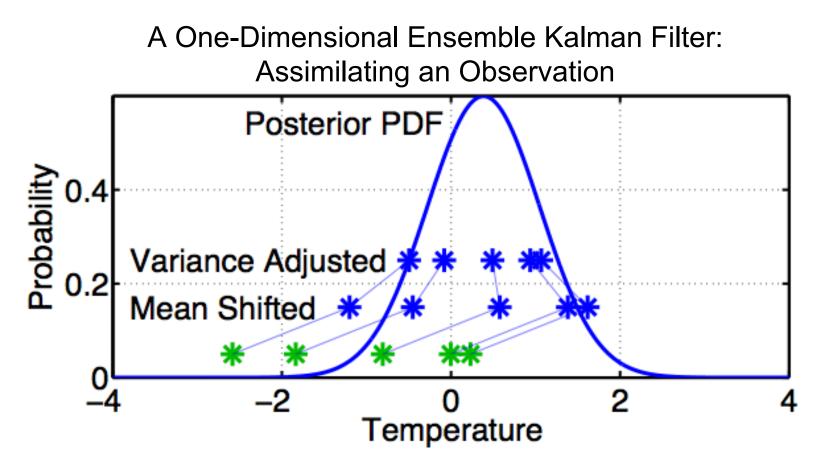




First, 'shift' the ensemble to have the exact mean of the posterior.







First, 'shift' the ensemble to have the exact mean of the posterior. Second, linearly contract to have the exact variance of the posterior. Sample statistics are identical to Kalman filter.



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UQ from an Ensemble Kalman Filter

(Ensemble) KF optimal for linear model, gaussian likelihood.
In KF, only mean and variance have meaning.

Variance defines the uncertainty.

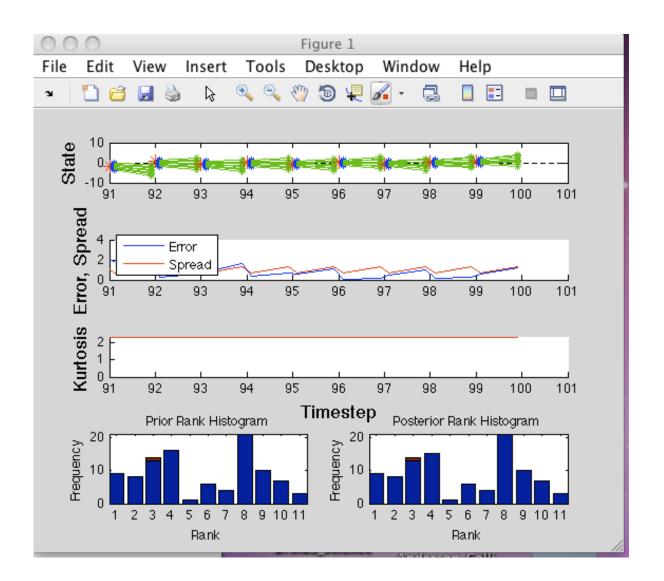
- Ensemble allows computation of many other statistics.
- What do they mean? Not entirely clear.
- Example: Kurtosis. Completely constrained by initial ensemble. It is problem specific whether this is even defined!

(See example set 1 in Appendix)



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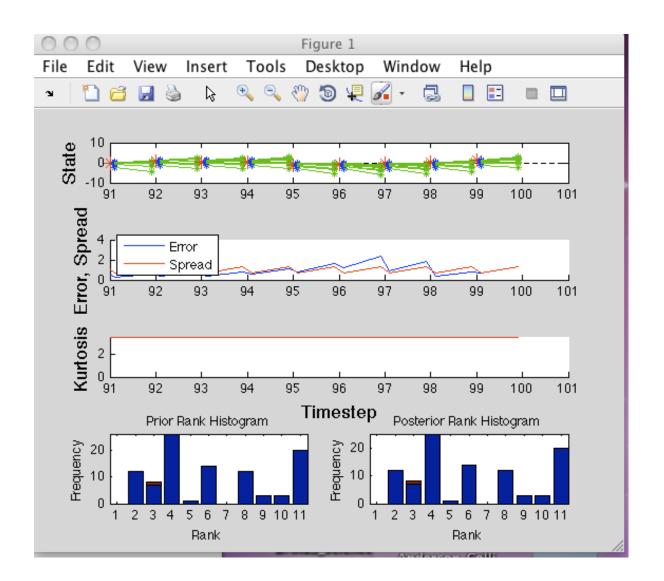






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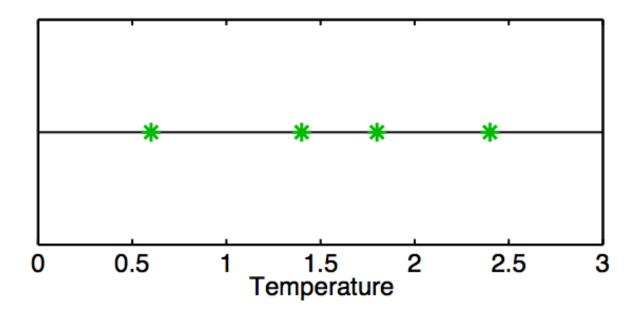




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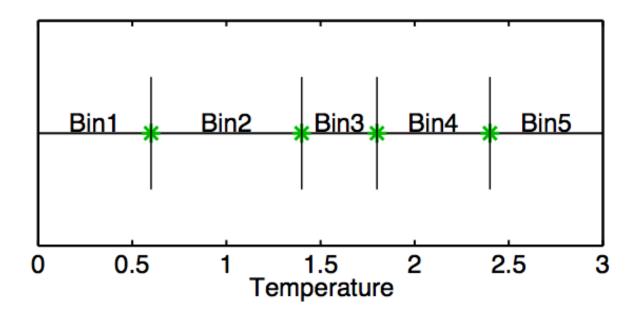
Draw 5 values from a real-valued distribution. Call the first 4 'ensemble members'.







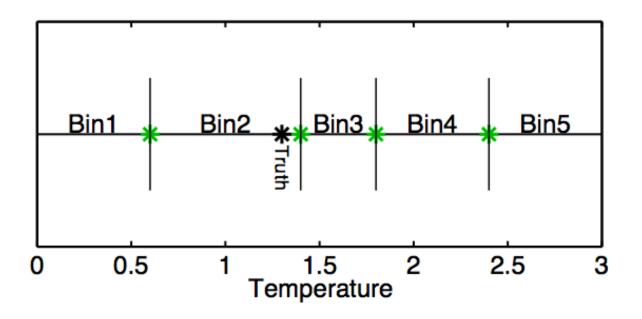
These partition the real line into 5 bins.





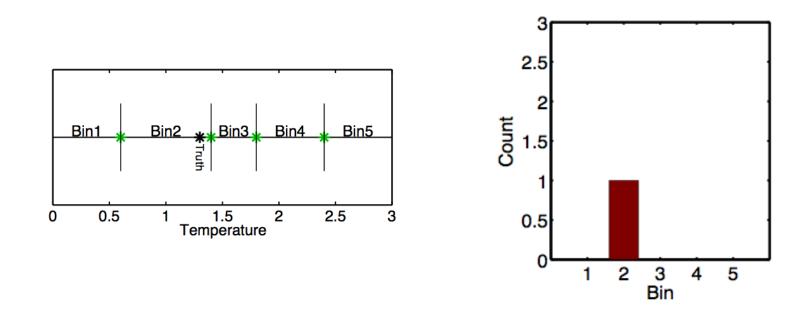


Call the 5th draw the 'truth'. 1/5 chance that this is in any given bin.



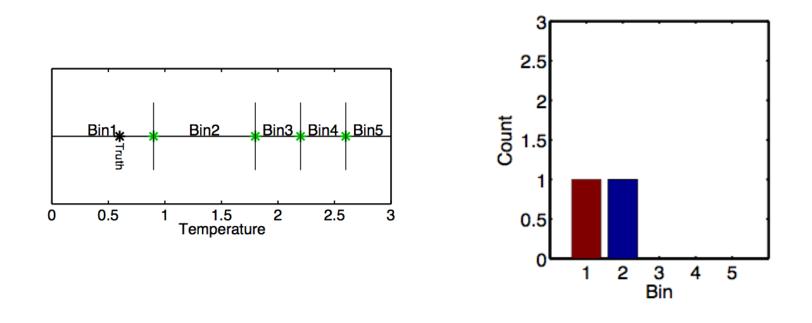






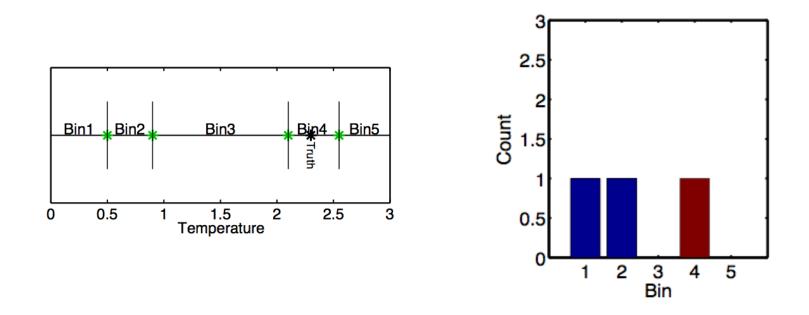






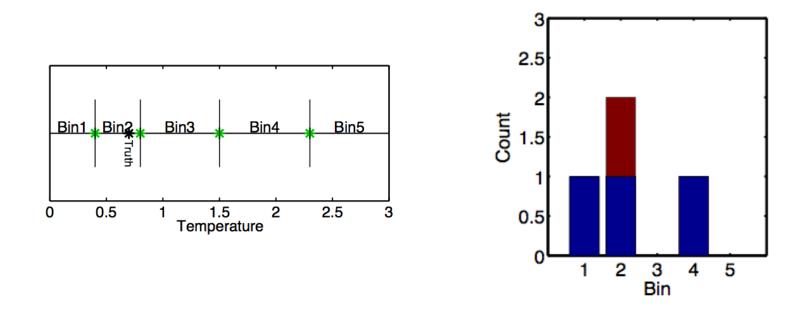








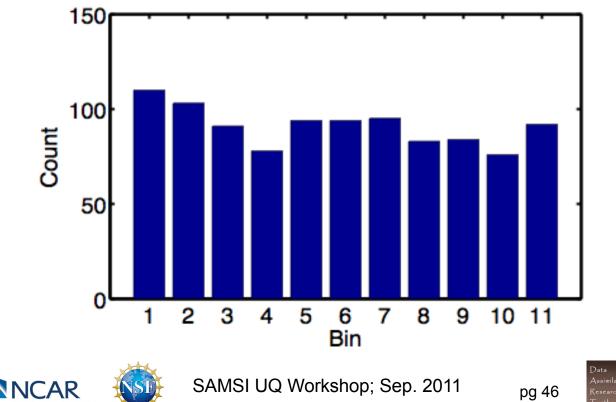






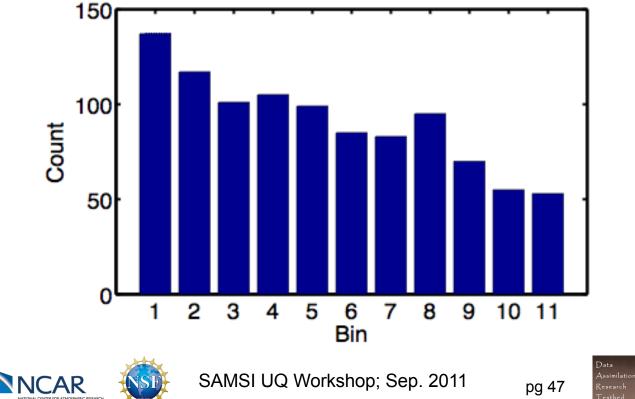


Rank histograms for good ensembles should be uniform (caveat sampling noise). Want truth to look like random draw from ensemble.



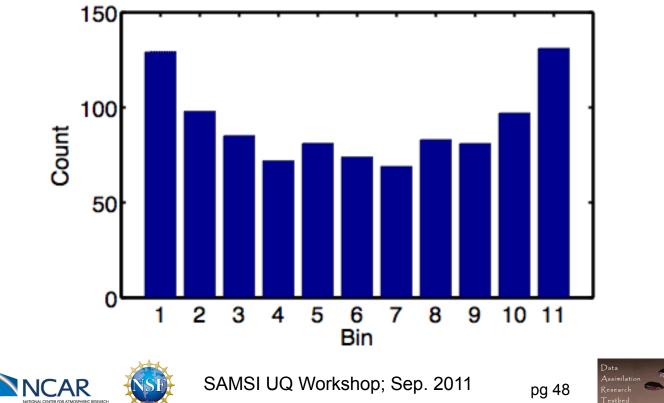


A biased ensemble leads to skewed histograms.





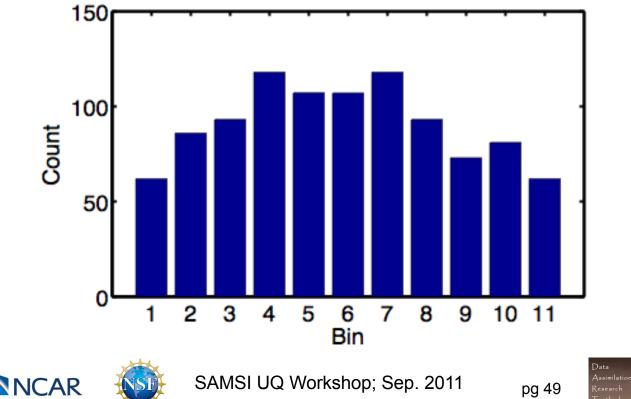
An ensemble with too little spread gives a u-shape. This is the most common behavior for geophysics.





An ensemble with too much spread is peaked in the center.

(See example set 1 in appendix)



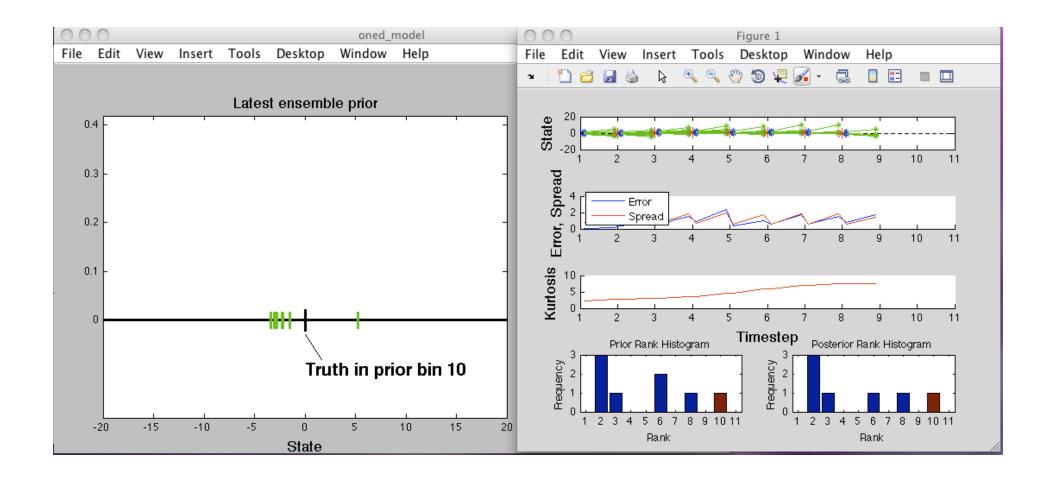


Back to 1-Dimensional (Ensemble) Kalman Filter

- All bets are off when model is nonlinear, likelihood nongaussian.
- Must assess quality of estimates for each case.
- Example: A weakly-nonlinear 1D model: dx/dt = x + 0.4|x|x (See example set 2 in appendix)
- Ensemble statistics can become degenerate.
- Nonlinear ensemble filters may address this but beyond scope for today.

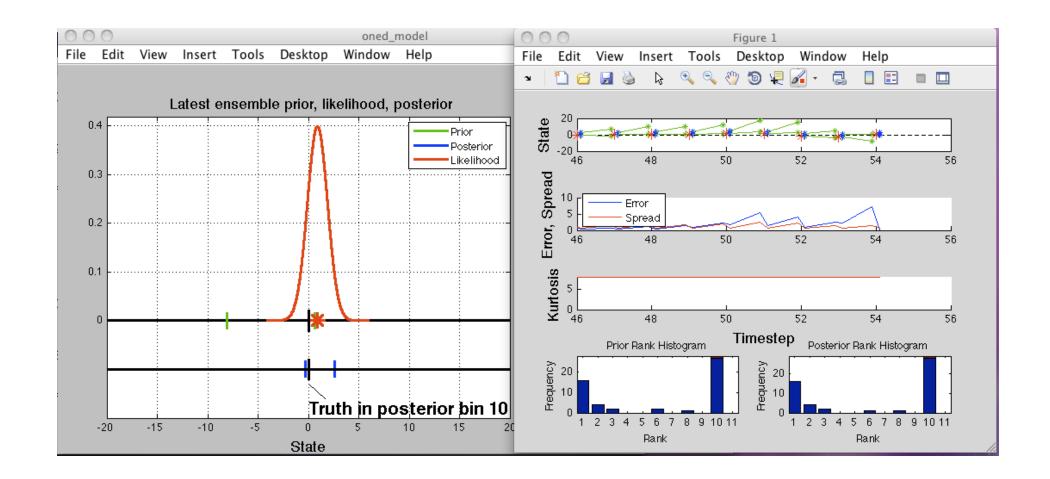














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Ensemble Kalman Filter with Model Error

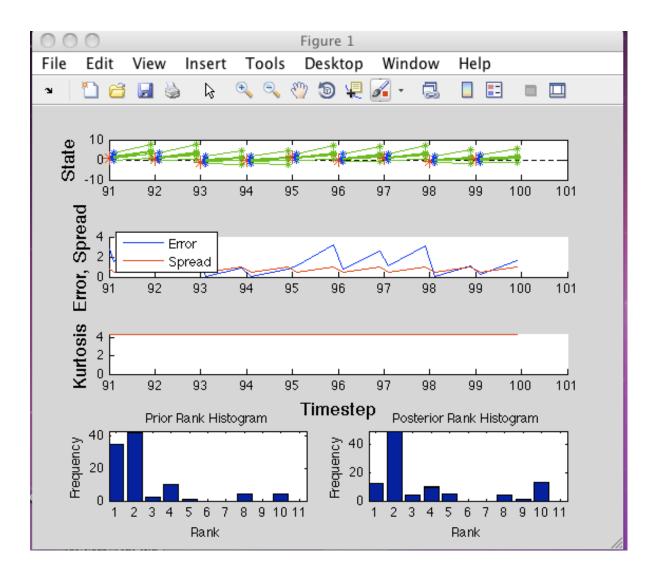
- Welcome to the real world. Models aren't perfect.
- ➢ Model (and other) error can corrupt ensemble mean and variance.

(See example set 3 part 1 in appendix)

- Adapt to this by increasing uncertainty in prior.
- Inflation is common method for ensembles.



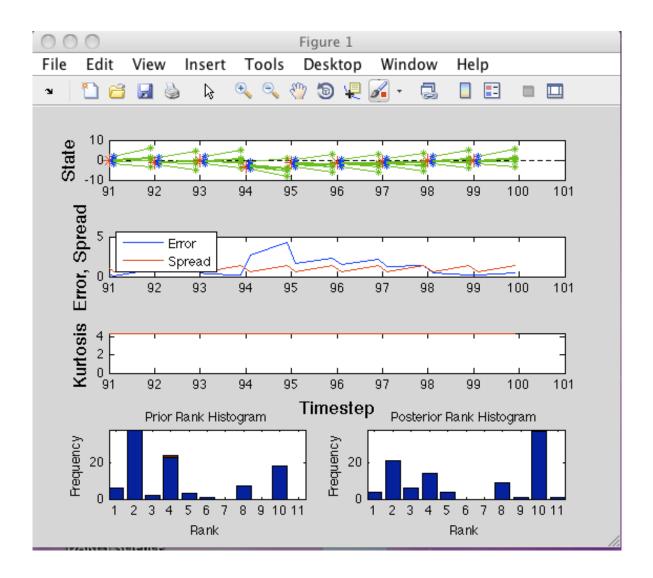






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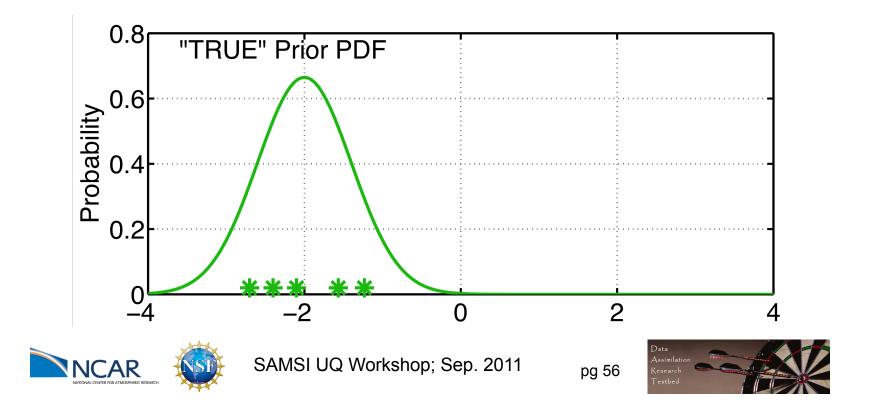


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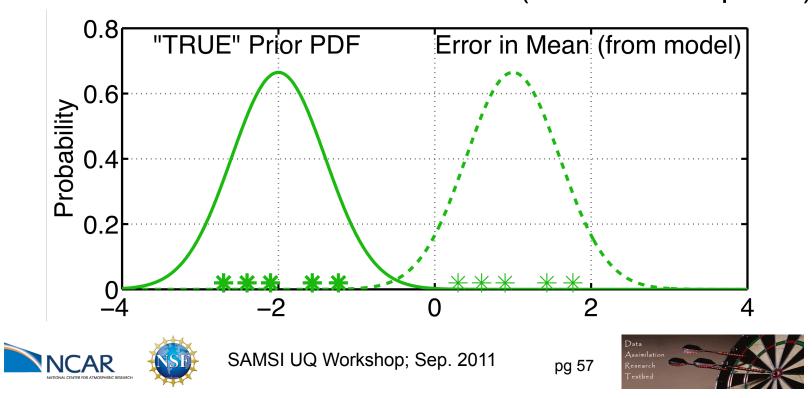
Dealing with systematic error: Variance Inflation

Observations + physical system => 'true' distribution.



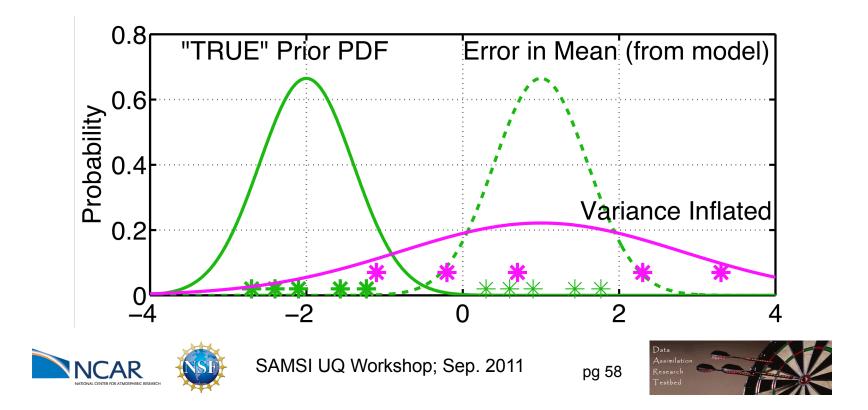
Dealing with systematic error: Variance Inflation

Observations + physical system => 'true' distribution. Model bias (and other errors) can shift actual prior. Prior ensemble is too certain (needs more spread).



Dealing with systematic error: Variance Inflation

Naïve solution: increase the spread in the prior. Give more weight to the observation, less to prior.



Ensemble Kalman Filter with Model Error

- Errors in model (and other things) result in too much certainty in prior.
- Can use inflation to correct for this.
- Generally tuned on dependent data set.
- Adaptive methods exist but must be calibrated. (See example set 3 part 2 in appendix)
- Uncertainty errors depend on forecast length.





Multivariate Ensemble Kalman Filter

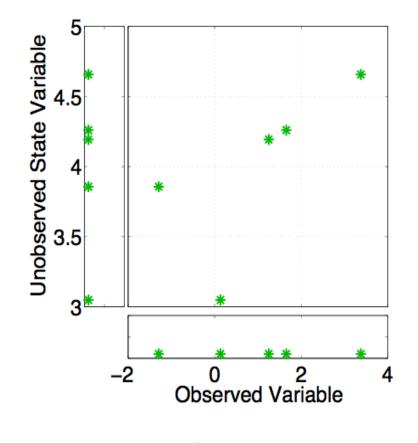
So far, we have an observation likelihood for single variable.

Suppose the model prior has additional variables.

Use linear regression to update additional variables.







Assume that all we know is prior joint distribution.

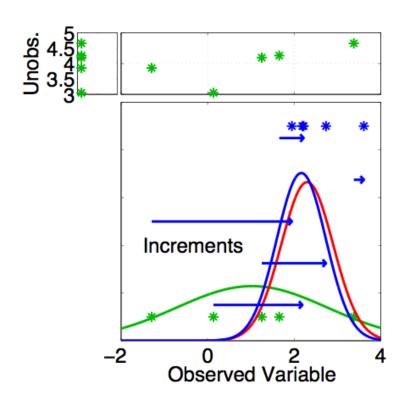
One variable is observed.

What should happen to the unobserved variable?



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Assume that all we know is prior joint distribution.

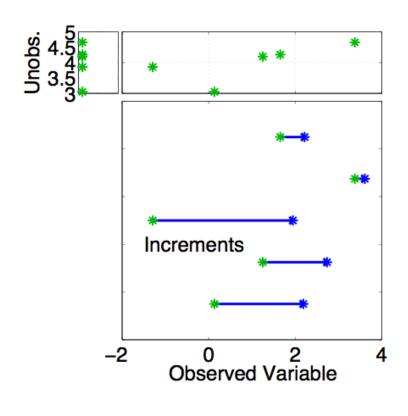
One variable is observed.

Compute increments for prior ensemble members of observed variable.



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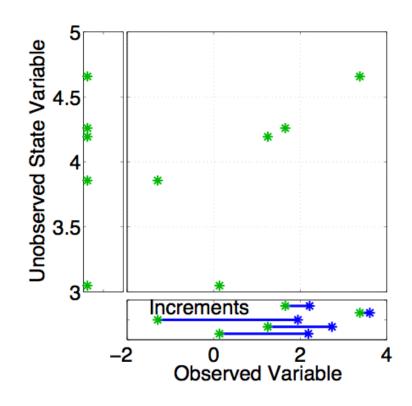
Assume that all we know is prior joint distribution.

One variable is observed.

Using only increments guarantees that if observation had no impact on observed variable, unobserved variable is unchanged (highly desirable).







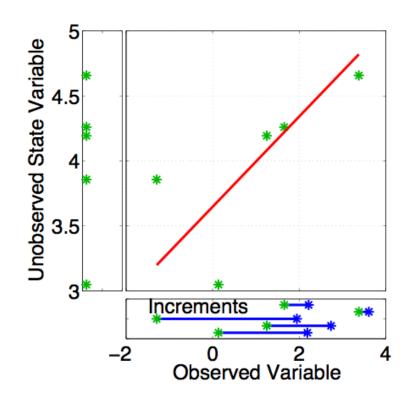
Assume that all we know is prior joint distribution.

How should the unobserved variable be impacted? First choice: least squares. Equivalent to linear regression. Same as assuming binormal prior.



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Have joint prior distribution of two variables.

How should the unobserved variable be impacted?

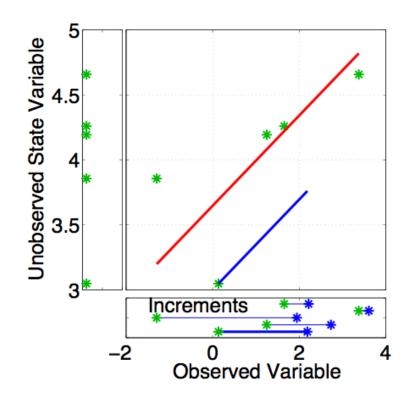
First choice: least squares.

Begin by finding <u>least squares</u> <u>fit.</u>



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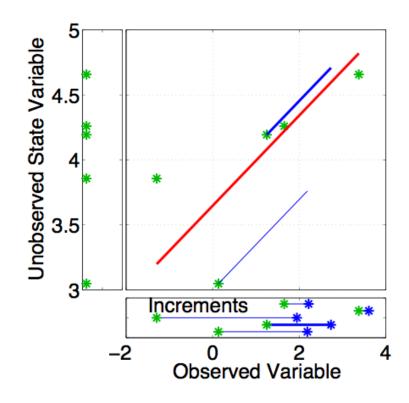


Have joint prior distribution of two variables.

Next, regress the observed variable increments onto increments for the unobserved variable.







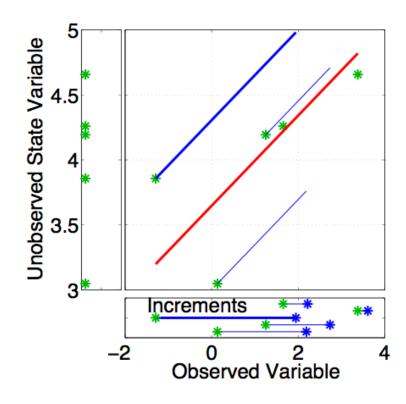
Have joint prior distribution of two variables.

Next, regress the observed variable increments onto increments for the unobserved variable.

Equivalent to first finding image of increment in joint space.





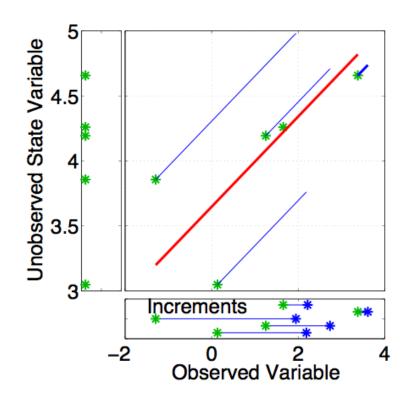


Have joint prior distribution of two variables.

Next, regress the observed variable increments onto increments for the unobserved variable.





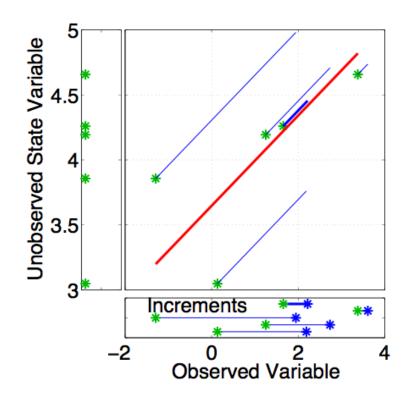


Have joint prior distribution of two variables.

Next, regress the observed variable increments onto increments for the unobserved variable.





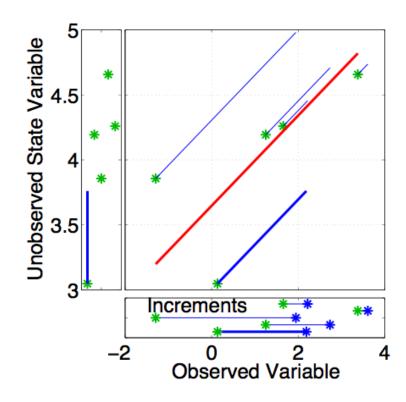


Have joint prior distribution of two variables.

Next, regress the observed variable increments onto increments for the unobserved variable.







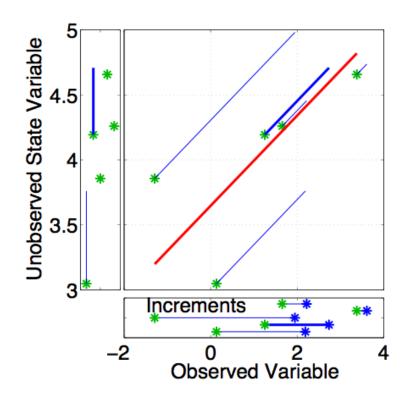
Have joint prior distribution of two variables.

Regression: Equivalent to first finding image of increment in joint space.

Then projecting from joint space onto unobserved priors.







Have joint prior distribution of two variables.

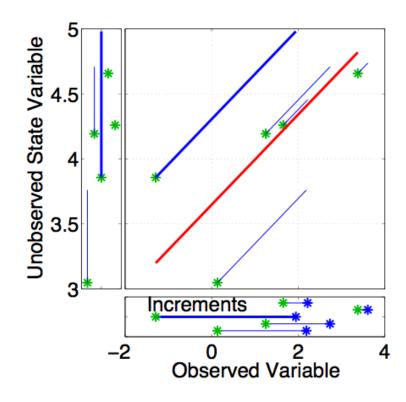
Regression: Equivalent to first finding image of increment in joint space.

Then projecting from joint space onto unobserved priors.





Ensemble filters: Updating additional prior state variables



Have joint prior distribution of two variables.

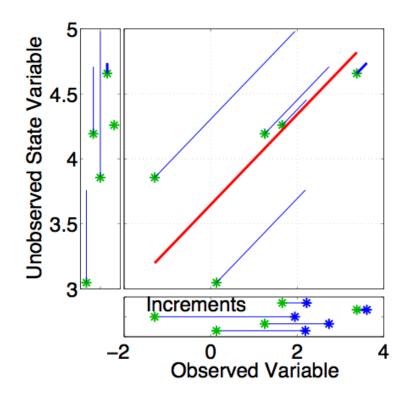
Regression: Equivalent to first finding image of increment in joint space.

Then projecting from joint space onto unobserved priors.





Ensemble filters: Updating additional prior state variables



Have joint prior distribution of two variables.

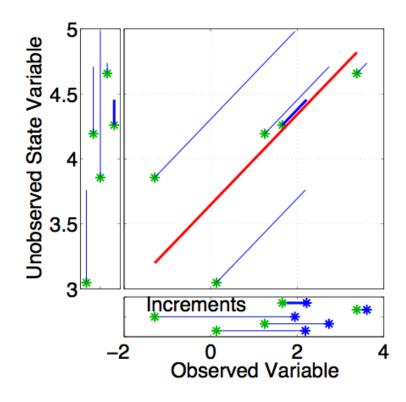
Regression: Equivalent to first finding image of increment in joint space.

Then projecting from joint space onto unobserved priors.





Ensemble filters: Updating additional prior state variables



Have joint prior distribution of two variables.

Regression: Equivalent to first finding image of increment in joint space.

Then projecting from joint space onto unobserved priors.





The Lorenz63 3-Variable 'Chaotic' Model

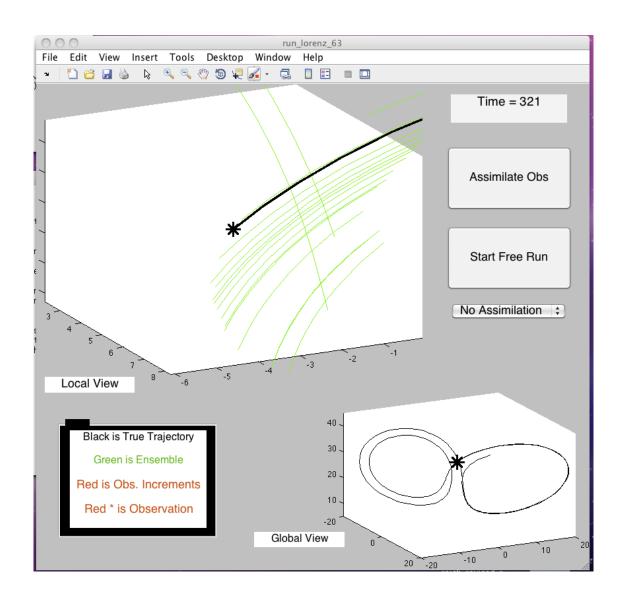
≻Two lobed attractor.

- ≻State orbits lobes, switches lobes sporadically.
- ≻Nonlinear if observations are sparse.
- ≻Qualitative uncertainty information available from ensemble.
- ≻Nongaussian distributions are sampled.
- ➤Quantitative uncertainty requires careful calibration.

(See Example 4 in the appendix)

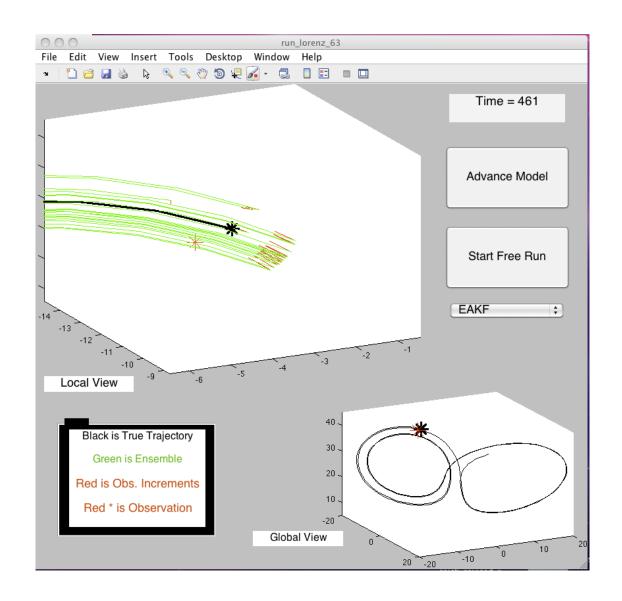
















Ensemble Kalman Filters with Large Models

➤Kalman filter is too costly (time and storage) for large models.

- ➤Ensemble filters can be used for very large applications.
- Sampling error when computing covariances may become large.

Spurious correlations between unrelated observations and state variables contaminate results.

>Leads to further uncertainty underestimates.

➤Can be compensated by a priori limits on what state variables are impacted by an observation. Called 'localization'.

≻Requires even more calibration.





The Lorenz96 40-Variable 'Chaotic' Model

- Something like weather systems around a latitude circle.
- ≻Has spurious correlations for small ensembles.
- ➢Basic implementation of ensemble filter leads to bad uncertainty estimates.
- ➢Both inflation and localization can improve performance.
- ≻Quantitative uncertainty requires careful calibration.

(See Example 5 in the appendix)





Uncertainty and Ensemble Kalman Filters: Conclusions

 \geq Ensemble KF variance is exact quantification of uncertainty when things are linear, gaussian, perfect, and ensemble is large enough.

➤Too little uncertainty nearly universal for geophysical applications.

≻Uncertainty errors are a function of forecast lead time.

>Adaptive algorithms like inflation improve uncertainty estimates.

≻Quantitative uncertainty estimates require careful calibration.

➢Out of sample estimates (climate change) are tenuous.





This is an outline of the live matlab demos from NCAR's DART system that were used to support this tutorial. The matlab code is available by checking out the DART system from NCAR at:

http://www2.image.ucar.edu/forms/dart-software-download

The example scripts can be found in the DART directory: DART_LAB/matlab

Some of the features of the GUI's for the exercises do not work due to bugs in Matlab releases 2010a and 2010b but should work in earlier and later releases. Many more features of these matlab scripts are described in the tutorial files in the DART directory:

DART_LAB/presentation

These presentations are a complement and extension of things presented in this tutorial.



DART-LAB Tutorial -- June 09



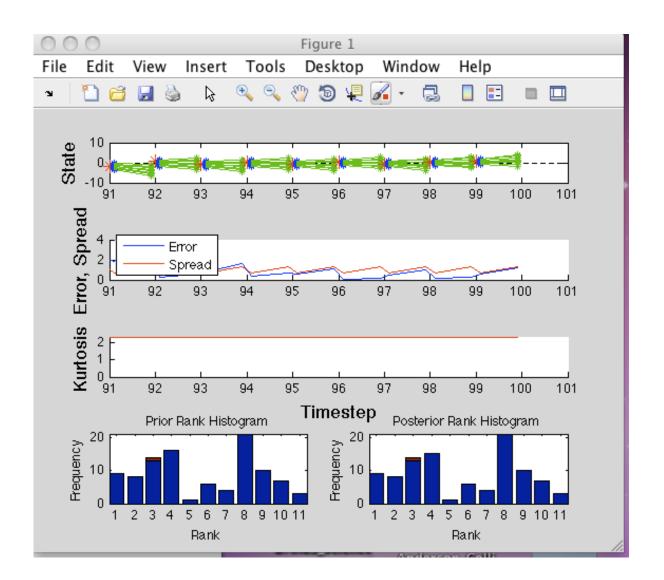
Example Set 1: Use matlab script oned_model Change the "Ens. Size" to 10 Select "Start Free Run"

Observe that the spread curve in the second diagnostic panel quickly converges to have the same prior and posterior values. The rms error varies with time due to observation error, but its time mean should be the same as the spread.

Restart this exercise several times to see that the kurtosis curve in the third panel also converges immediately, but to values that change with every randomly selected initial ensemble. Also note the behavior of the rank histograms for this case. They may not be uniform even though the ensemble filter is the optimal solution. In fact, they can be made arbitrarily nonuniform by the selection of the initial conditions.

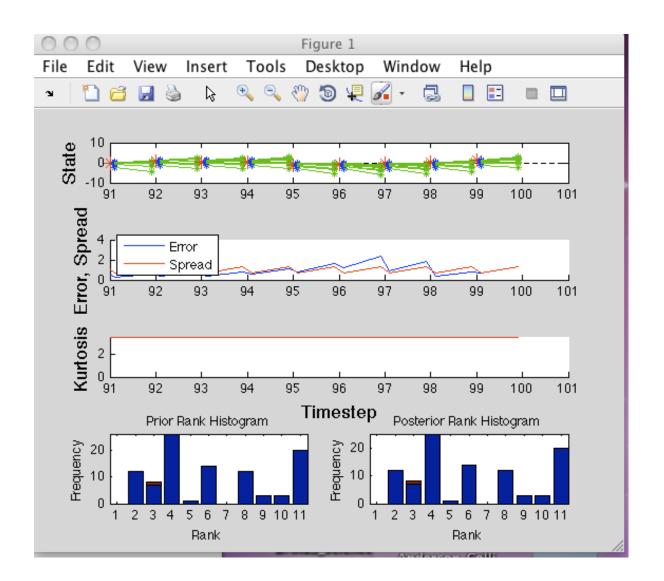














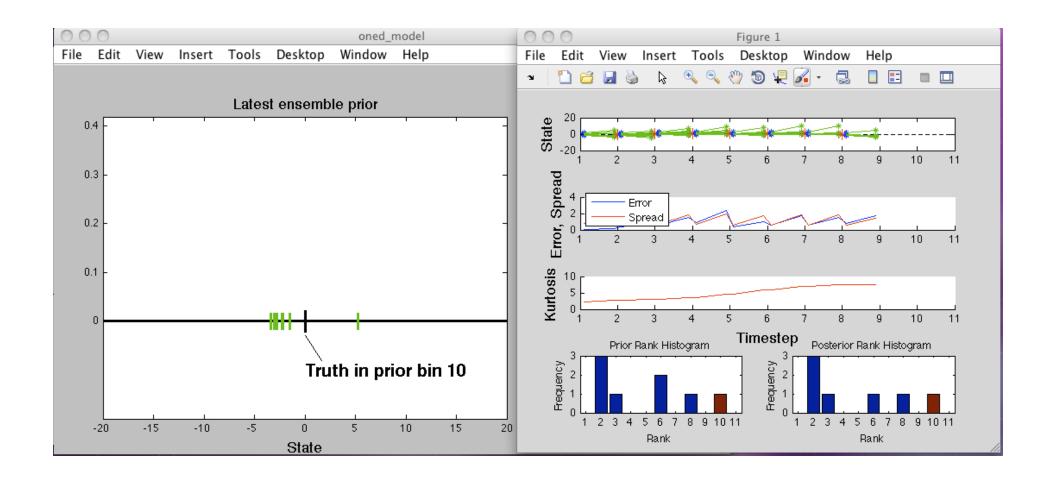


Example Set 2: Use matlab script oned_model Change the "Ens. Size" to 10 Change "Nonlin. A" to 0.4 Select "Start Free Run"

This examines what happens when the model is nonlinear. Observe the evolution of the rms error and spread in the second panel and the kurtosis in the third panel. Also observe the ensemble prior distribution (green) in the first panel and in the separate plot on the control panel (where you changed the ensemble size). In many cases, the ensemble will become degenerate with all but one of the ensemble members collapsing while the other stays separate. Observe the impact on the rank histograms as this evolves.

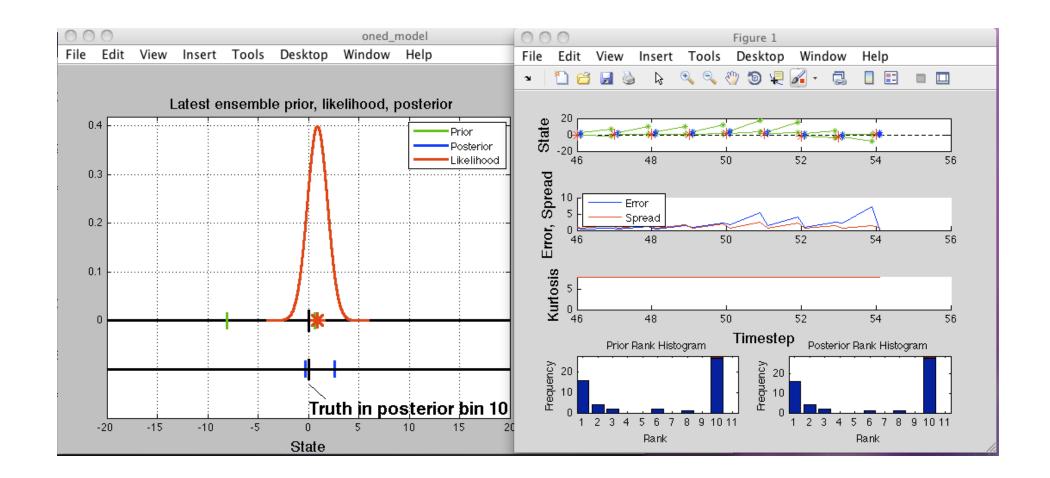














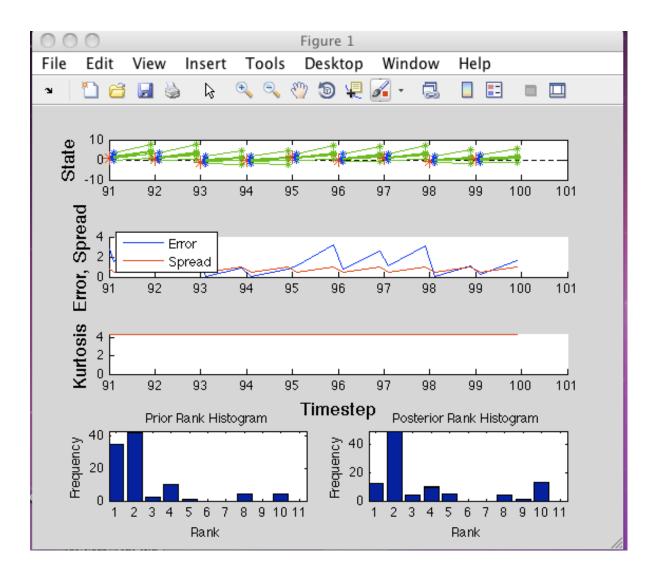


Example Set 3: Use matlab script oned_model Change the "Ens. Size" to 10 Change "Model Bias" to 1 Select "Start Free Run"

This examines what happens when the model used to assimilate observations is biased compared to the model that generated the observations (an imperfect model study). The model for the assimilation adds 1 to what the perfect model would do at each time step. Observe the evolution of the rms error and spread in the second panel and the rank histograms. Now, try adding some uncertainty to the prior ensemble using inflation by setting "Inflation" to 1.5. This should impact all aspects of the assimilation.

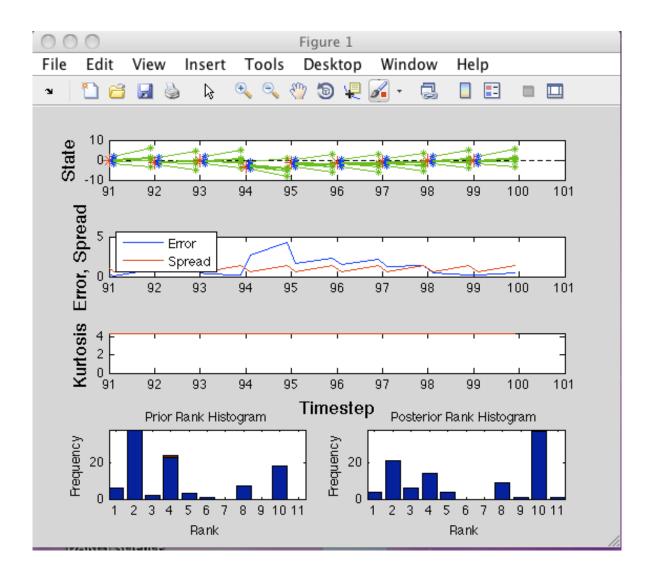














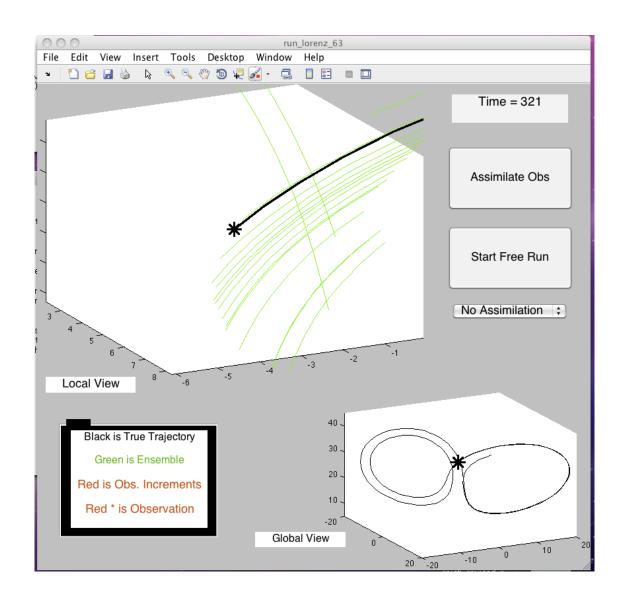


Example Set 4: Use matlab script run_lorenz_63 Select "Start Free Run"

This run without assimilation shows how uncertainty grows in an ensemble forecast in the 3 variable Lorenz model. After a few trips around the attractor select "Stop Free Run". Then turn on an ensemble filter assimilation by changing "No Assimilation " to "EAKF". Continue the run by selecting "Start Free Run" and observe how the ensemble evolves. You can stop the advance again and use the matlab rotation feature to study the structure of the ensemble.

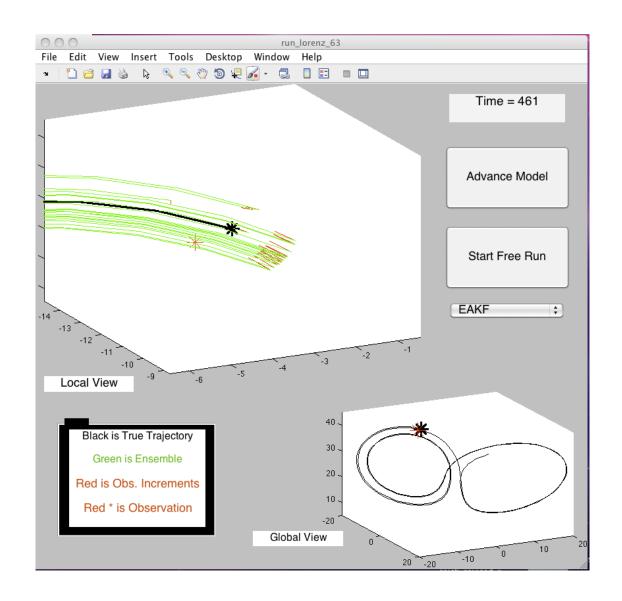
















Example Set 5: Use matlab script run_lorenz_96 Select "Start Free Run"

This run without assimilation shows how uncertainty grows in an ensemble forecast in the 40 variable model. After about time=40, select "Stop Free Run" and start an assimilation by changing "No Assimilation" to "EAKF" and selecting "Start Free Run". Exploring the use of a localization of 0.2 and/or an inflation of 1.1 helps to show how estimates of uncertainty can be improved by modifying the base ensemble Kalman filter.



