Prediction Uncertainty in Analysis of a Computer Experiment: Fast Bayesian Inference

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Outline Arctic Sea Ice Example Computer Codes in General Bayesian Formulation Fast Bayesian Inference (FBI) Conclusions

Outline of Topics

Arctic Sea Ice Example

Computer Codes in General

Bayesian Formulation

Fast Bayesian Inference (FBI)

Conclusions



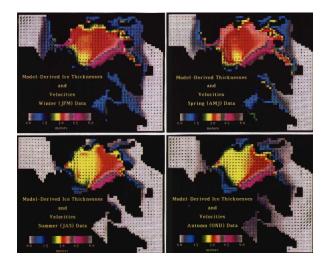
Arctic Sea Ice Example

Chapman, Welch, Bowman, Sacks and Walsh, "Arctic Sea Ice Variability: Model Sensitivities and a Multidecadal Simulation", *Journal of Geophysical Research*, 99C (1994), 919–935



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Arctic Sea Ice: Code Output





Arctic Sea Ice: Code Input and Output Variables

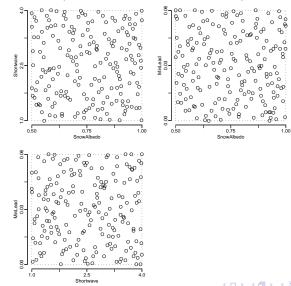
• 13 input variables:

SnowAlbedo, Shortwave, MinLead, OceanicHeat, Snowfall, Longwave, AtmosDrag, SensHeat, LatentHeat, LogIceStr, OceanicDrag, OpenAlbedo, IceAlbedo

- 4 output variables: IceMass, IceArea, IceVelocity, RangeOfArea
- 157 good runs of the code (out of 191 attempts)

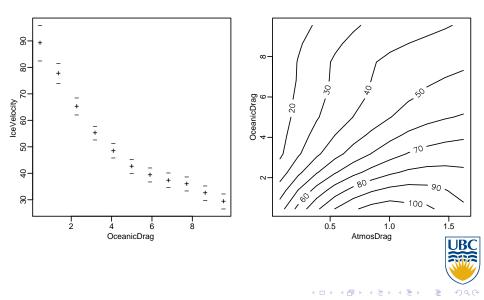


Arctic Sea Ice: Experimental Design (First 3 Input Variables)



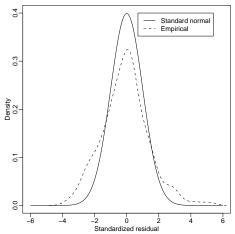


Arctic Sea Ice: Sensitivity Analysis for Ice Velocity



Arctic Sea Ice: The Inference Problem Illustrated

Standardized Residuals = $(y(\mathbf{x}) - \hat{y}(\mathbf{x}))/se(\hat{y}(\mathbf{x}))$ from 10-fold cross-validation (RangeOfArea):





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Computer Codes in General

Computer code $y(\mathbf{x})$

- $\mathbf{x} = (x_1, \dots, x_d)$ is a vector of d input variables
- y is a (scalar) output variable



Statistical Approximating Model

Approximation / prediction / emulation of $y(\mathbf{x})$ is the "engine" of analysis of computer experiments:

- To replace the computer model in future with a fast surrogate
- Sensitivity analysis
- Visualization
- Optimization
- Code calibration and assessment



The unknown function $y(\mathbf{x})$ of vector-valued \mathbf{x} is treated as if it were a *realization of a random function* (RF) or *Gaussian Stochastic Process* (GaSP), $Z(\mathbf{x})$:

• Z(x) has a Gaussian (normal) distribution at any x.



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$$Var(Z(\mathbf{x})) = \sigma_Z^2$$
, i.e., is constant.



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- $Z(\mathbf{x})$ has mean β (can be generalized to regression terms).
- $Var(Z(\mathbf{x})) = \sigma_Z^2$, i.e., is constant.
- At two input vectors, **x** and **x**', the corresponding Z's are correlated,

$$\operatorname{Corr}(Z(\mathbf{x}), Z(\mathbf{x}')) \equiv R(\mathbf{x}, \mathbf{x}'),$$

where R is a correlation function known up to various parameters.

 $Corr(Z(\mathbf{x}), Z(\mathbf{x}'))$ is key to the random function model.



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Power-Exponential Correlation

In one dimension: $R(x, x') = e^{-\theta |x-x'|^p}$.

- $\theta \ge 0$ controls the activity of the function.
 - Larger θ gives smaller correlation, i.e., $Z(\mathbf{x})$ and $Z(\mathbf{x}')$ are less related and the function is more complex.
- $p \in [1,2]$ affects the smoothness of the random function.
 - p = 2 (Gaussian or squared exponential correlation function) gives very smooth realizations (good for approximating functions with many derivatives).
 - *p* = 1 gives much rougher realizations (good for continuous but non-differentiable functions).

Much used in practice.



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Statistical Approximator/Predictor/Emulator (Kriging)

n runs, input vectors $\mathbf{x}^{(1)}, \dots, \mathbf{x}^{(n)}$, and output values \mathbf{y} . Approximator of $y(\mathbf{x})$ at any \mathbf{x} :

$$\hat{y}(\mathbf{x}) = \hat{\beta} + \mathbf{r}(\mathbf{x})^{T}(\mathbf{x})\mathbf{R}^{-1}(\mathbf{y} - \mathbf{1}\hat{\beta}),$$

where

- $\hat{\beta} = \mathbf{1}^T \mathbf{R}^{-1} \mathbf{y} / \mathbf{1}^T \mathbf{R}^{-1} \mathbf{1}$
- 1 is an *n* × 1 vector of 1's
- **R** is an $n \times n$ correlation matrix, with element i, j given by $R(\mathbf{x}^{(i)}, \mathbf{x}^{(j)})$
- $\mathbf{r}(\mathbf{x})$ is an $n \times 1$ vector of correlations with element *i* given by $R(\mathbf{x}^{(i)}, \mathbf{x})$



Standard Error of the Approximator

Standard error of $\hat{y}(\mathbf{x})$:

$$\mathsf{se}(\hat{y}(\mathbf{x})) = \sigma^2 \left[1 - \mathbf{r}(\mathbf{x})' \mathbf{R}^{-1} \mathbf{r}(\mathbf{x}) + \frac{\left(1 - \mathbf{1}' \mathbf{R}^{-1} \mathbf{r}(\mathbf{x})\right)^2}{\mathbf{1}' \mathbf{R}^{-1} \mathbf{1}} \right],$$

where r(x), R, and 1 were defined on the previous slide. We have to estimate the correlation parameters

$$\theta_1,\ldots,\theta_d,\ p_1,\ldots,p_d,$$

e.g., by maximum likelihood.

THE PROBLEM: If we just PLUG-IN the estimates of the correlation parameters, the above $se(\hat{y}(x))$ ignores this source of uncertainty and is too small.



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Bayesian Formulation

In principle, Bayes can deal with all sources of uncertainty. Let θ denote all the correlation parameters. Posterior for θ given **y** from Bayes rule:

 $p(\theta \mid \mathbf{y}) \propto \pi(\theta) L(\mathbf{y} \mid \theta)$

- $\pi(\theta)$ is the prior for heta
- $L(\mathbf{y} \mid \boldsymbol{\theta})$ is the multivariate normal likelihood arising from the RF model.

The predictive distribution for $y(\mathbf{x})$ is

$$p(y(\mathbf{x} | \mathbf{y})) = \int p(y(\mathbf{x}) | \mathbf{y}, \theta) p(\theta | \mathbf{y}) d\theta.$$



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Objections

I want to be a Bayesian BUT...

1. I'll have to choose a prior, i.e., $\pi(\theta)$.



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Objections

I want to be a Bayesian BUT...

- 1. I'll have to choose a prior, i.e., $\pi(\theta)$.
- 2. I'll have to run MCMC.

Relax. You won't have to!





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FBI: Main Results

- 1. Results:
 - Even the (computationally) cheapest Bayesian implementation beats the plug-in method in terms of validity of uncertainty measures.
- 2. Methods:
 - Plug-in (Kriging)
 - Fast Bayesian Implementation (FBI)
 - Fast to implement
 - Fast to execute



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FBI Algorithm

Only implemented for the Gaussian correlation function, $R(x, x') = e^{-\theta |x-x'|^2}$.

- 1. Use the $\theta_j^* = \log \theta_j$ transformation.
- 2. Maximize the log likelihood to get the MLE of θ^* and the HESSIAN at the MLE.
- 3. Sample θ^* from $N(\text{MLE}, -\text{HESSIAN}^{-1})$ to get:

•
$$\hat{y}_{\text{FBI}} = \mathbf{E}_{\boldsymbol{\theta}^*}(\hat{y}_{\text{Plug-in}})$$

- $\operatorname{Var}(\hat{y}_{\operatorname{FBI}}) = \operatorname{Var}_{\theta^*}(\hat{y}_{\operatorname{Plug-in}}) + \operatorname{E}_{\theta^*}(\operatorname{Var}(\hat{y}_{\operatorname{Plug-in}}))$
- 4. Sampling is straightforward Monte Carlo or Quasi Monte Carlo (100 samples)! Not MCMC.



Simulations

- Generate y(x) at N values of x ∈ [0,1]^d as a realization of a Gaussian Random Function, Z(x), with Gaussian correlation function.
- Choose *n* of the **x** points as an experimental design, and use the corresponding realized $y(\mathbf{x}) = Z(\mathbf{x})$ values to fit the RF model.
- Predict at the remaining N n test points.
- Compute actual coverage probabilities of the predictions (should equal the nominal coverage probability).

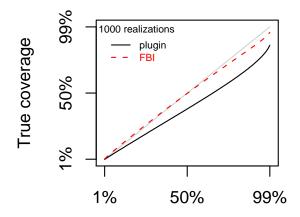
As $y(\mathbf{x})$ is sampled from the assumed statistical model, uncertainty estimates should be EXACT, leading to correct coverage probabilities.



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Simulation Results for d = 1, n = 10, and $\theta = 2$

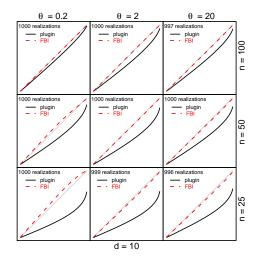


Nominal coverage



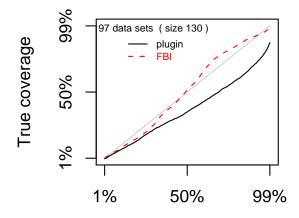
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Simulation Results for 10-dimensional x





Arctic Sea Ice: d = 13 and n = 130



Nominal coverage



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Summary

- 1. Methods:
 - Plug-in
 - Fast Bayesian Implementation (FBI) (extends the thesis of Karuri, 2005)
 - Fast to implement
 - Fast to execute
- 2. Results:
 - Even the (computationally) cheapest Bayesian implementation beats the plug-in method (consistent with Currin et al. 1991)
 - FBI has many approximations but IT WORKS!



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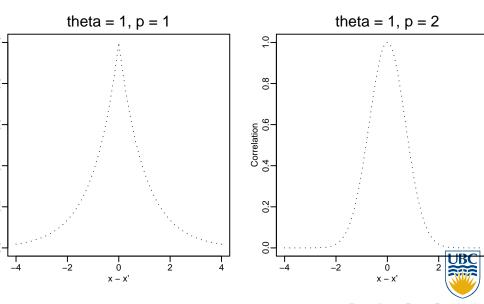
The End—Thank You

References on previous page:

- Currin, C., Mitchell, T., Morris, M., and Ylvisaker, D. (1991). "Bayesian Prediction of Deterministic Functions With Applications to the Design and Analysis of Computer Experiments." *Journal of the American Statistical Association*, 86, 953-963.
- Karuri, S. (2005). "Integration in Computer Experiments and Bayesian Analysis." Ph.D Thesis, University of Waterloo.



Power-Exponential Correlation



Correlation in Several Dimensions

With *d*-dimensional **x**, i.e.,

$$\mathbf{x}=(x_1,\ldots,x_d),$$

we simply use a product correlation rule, i.e.,

$$R(\mathbf{x},\mathbf{x}') = \prod_j R_j(x_j,x_j') = \prod_j e^{- heta_j|x_j-x_j'|^{p_j}}.$$

