Inference for spatial fields

Douglas Nychka,

www.image.ucar.edu/~nychka

- A spatial model and Kriging
- Kriging = Penalized least squares
- The Bayes connection
- Identifying a covariance function





The additive model

Given n pairs of observations (x_i, y_i) , i = 1, ..., n

$$y_i = g(x_i) + \epsilon_i$$

 ϵ_i 's are random errors.

Assume that g is a realization of a Gaussian process. and ϵ are $MN(0, \sigma^2 I)$

Formulating a statistical model for g makes a very big difference in how we solve the problem.

A Normal World

We assume that g(x) is a Gaussian process,

$$\rho k(\mathbf{x}, \mathbf{x}') = COV(g(\mathbf{x}), g(\mathbf{x}'))$$

For the moment assume that E(g(x)) = 0.

(A Gaussian process \equiv any subset of the field locations has a multivariate normal distribution.)

We know what we need to do!

If we know k we know how to make a prediction at x!

$$\hat{g}(x) = E[g(x)|data]$$

i.e. Just use the conditional multivariate normal distribution.

A review of the conditional normal

 $u \sim N(0, \Sigma)$

and

$$u = \left(egin{array}{c} u_1 \ u_2 \end{array}
ight) \; \Sigma = \left(egin{array}{c} \Sigma_{11}, \Sigma_{12} \ \Sigma_{21}, \Sigma_{22} \end{array}
ight)$$

$$[u_2|u_1] = N(\Sigma_{2,1}\Sigma_{1,1}^{-1}u_1, \Sigma_{2,2} - \Sigma_{2,1}\Sigma_{1,1}^{-1}\Sigma_{1,2})$$

Our application is

 $u_1 = y$ (the Data)

and

 $u_2 = (g(x_1, ...g(x_N)))$ a vector of function values where we would like to predict.

The Kriging weights

Conditional distribution of g given the data y is Gaussian.

Conditional mean

$$\hat{g} = COV(g, y) [COV(y)]^{-1} y = Ay$$

rows of A are the Kriging weights.

Conditional variance

$$COV(\boldsymbol{g}, \boldsymbol{g}) - COV(\boldsymbol{g}, \boldsymbol{y}) \left[COV(\boldsymbol{y})\right]^{-1} COV(\boldsymbol{y}, \boldsymbol{g})$$

These two pieces characterize the entire conditional distribution

Kriging as a smoother

Suppose the errors are uncorrelated Normals with variance σ^2 .

$$\rho K = COV(g, y) = COV(g, g)$$
 and $COV(y) = (\rho K + \sigma^2 I)$

$$\hat{h} = \rho K (\rho K + \sigma^2 I)^{-1} y$$

$$= K(K + \lambda I)^{-1} y = A(\lambda) y$$

My geostatistics/BLUE overhead

For any covariance and any smoothing matrix (not just S above) we can easily derive the prediction variance.

Question find the minimum of

$$E\left[(g(\boldsymbol{x})-\widehat{g}(\boldsymbol{x}))^2\right]$$

over all choices of S. The answer: The Kriging weights ... or what we would do if we used the Gaussian process and the conditional distribution.

Folklore and intuition:

The spatial estimates are not very sensitive if one uses suboptimal weights, especially if the observations contain some measurement error.

It does matter for measures of uncertainty.

Kriging with a fixed part

Adding a fixed component

$$g(x) = \sum_{i} \phi_i(x) d_i + h(x)$$

d is fixed

h is a mean zero process with covariance, k.

The BLUE/Universal Kriging estimate is:

Find d by Generalized least squares

$$\hat{\boldsymbol{d}} = \left(T^T M^{-1} T\right)^{-1} T^T M^{-1} \boldsymbol{y}$$

"Krig" the residuals

$$\hat{h} = K(K + \lambda I)^{-1} (y - T\hat{d})$$

In general:

$$\widehat{g}(x) = \sum_{i} \phi_{i}(x)\widehat{d}_{i} + \sum_{j} k(x, x_{j})\widehat{c}_{j}$$

with

$$\hat{c} = (K + \lambda I)^{-1} (y - T\hat{d})$$

The connection to penalized least squares, splines and the smoothing parameter

Basis functions:

determined by the covariance function

Penalty function

K is based on the covariance.

The minimization criteria:

$$\min_{\boldsymbol{d},\boldsymbol{c}} \sum_{i=1}^{n} (\boldsymbol{y} - (T\boldsymbol{d} + K\boldsymbol{c})_i)^2 + \lambda \boldsymbol{c}^T K \boldsymbol{c}$$

The Kriging estimator is a spline with reproducing kernel k.

 λ is proportional to the measurement (nugget) variance

The Bayes connection

Bracket notation is very useful:

[Z] the density function for the random variable Z [Y|z] the conditional density function for the random variable Y given z.

$$[y|g]$$
 the likelihood for the data

$$[\boldsymbol{y}|g] \sim MN(\boldsymbol{g}, \sigma^2 I)$$

[g] the prior for g.

$$[g] \sim MN(0, \rho K)$$

Bayes Theorem: the posterior

$$[g|y] = \frac{[y|g][g]}{[y]} \sim [y|g][g]$$

The Posterior mode: where [g|y] has a maximum.

Maximizing:

[g|y] is the same as

Minimizing:

$$-2ln[g|y] = -2ln([y|g]) - 2ln([g]) + 2ln([y])$$

The posterior mode is the penalized least squares estimate where the penalty is equivalent to a prior!

This is true even we let "g" be the entire field, not just its values at the observations.

A causal example of identifying a covariance function

A useful form for k are isotropic correlations:

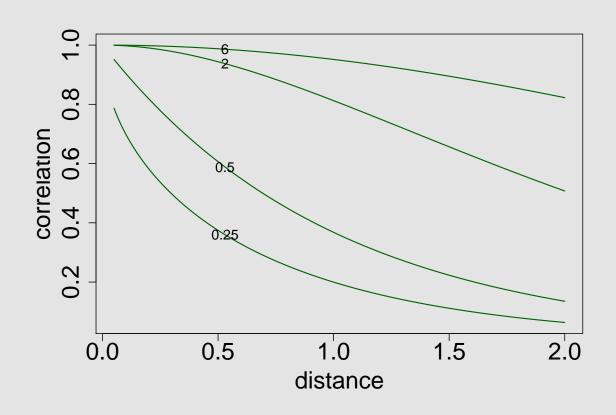
$$k(x, x') = \sigma(x)\sigma(x')\phi(||x - x'||)$$

The Matern class of covariances:

$$\phi(d) = \rho \psi_{\nu}(d/\theta))$$

 θ a range parameter, ν smoothness at 0. ψ_{ν} is an exponential for $\nu=1/2$ as $\nu\to\infty$ Gaussian.

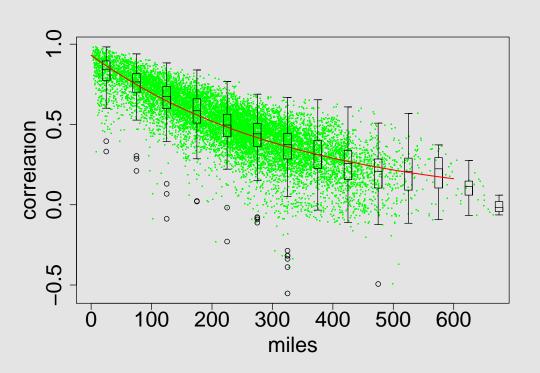
Matern family: the shape ν



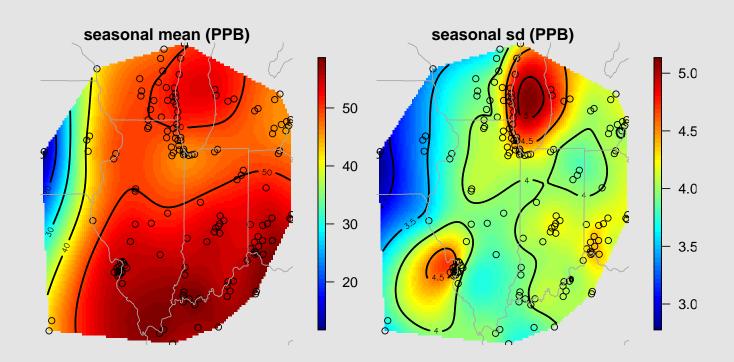
($\mathbf{m}^t h$ order thin plate spline in $\Re^d \nu = 2m - d!$)

Using the temporal information

In many cases spatial processes also have a temporal component. Here we take the 89 days over the "ozone season" and just find sample correlations among stations.



Mean and SD surfaces for 1987 ozone



Covariance model:

$$k(x, x') = \rho \sigma(x) \sigma(x') exp(-||x - x'||/\theta)$$

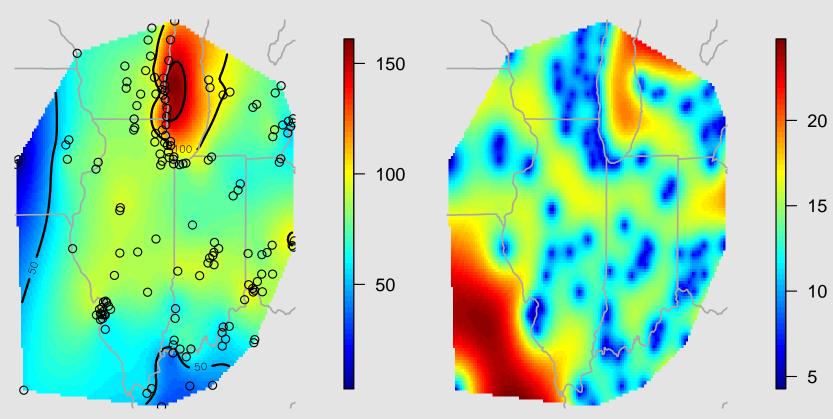
Mean model: $E(z(x)) = \mu(x)$

where μ is also a Gaussian spatial process.

Spatial estimate and uncertainty



Posterior standard deviation.



Summary

A spatial process model leads to a penalized least squares estimate

A spline = Kriging estimate = Bayesian posterior mode

For spatial estimators the basis functions are related to the covariance functions and can be identified from data